Three versions of the Short Alcohol Dependence Data Questionnaire (SADD) in Mexican population: a comparative analysis of psychometric properties and diagnostic accuracy

Alejandro Pérez-López,¹ Luis Villalobos-Gallegos,¹ María del Carmen Viveros Domínguez,² Javier Graue-Moreno,¹ Rodrigo Marín-Navarrete¹

Original article

ABSTRACT

Background

The Short Alcohol Dependence Data Questionnaire (SADD) has shown good reliability and validity in previous studies. In Mexico, although it is widely used in addiction treatment centers, little is known about its psychometric properties and diagnostic accuracy.

Objective

Hence, this study performed a Confirmatory Factorial Analysis (CFA) on three SADD versions (15, 14 and 12 items) and examined their operating characteristics.

Method

The sample included 570 individuals from the 30 Addiction Residential Centers localized in the central zone of Mexico.

Results

The three versions showed an internal consistency of >.90, fair goodness-of-fit, and significant correlations with the Mini International Neuropsychiatric Interview (MINI) 5.0 Alcohol dependence (AD) diagnostic criteria. The analysis of the operating characteristics revealed that each version accounted for 84-85% of the area under the curve (AUC).

Discussion and conclusion

The three SADD versions possess reliability and validity properties for the assessment of the alcohol dependence syndrome (ADS) and adequate diagnostic accuracy for the improvement of patients with AD in residential settings.

Key words: Alcohol dependence, SADD, psychometric properties, construct validity, diagnostic accuracy.

RESUMEN

Antecedentes

La Breve Escala de Dependencia al Alcohol (BEDA) ha mostrado confiabilidad y validez en estudios previos. En México, a pesar de que se ha utilizado ampliamente en programas de tratamiento para las adicciones, se conoce muy poco sobre sus propiedades psicométricas y su eficiencia diagnóstica.

Objetivo

Por lo tanto, el presente estudio realizó un Análisis Factorial Confirmatorio (AFC) en tres diferentes versiones de la BEDA (15, 14 y 12 reactivos) y examinó sus características operativas.

Método

La muestra incluyó 570 personas que recibían atención en centros de tratamiento residencial para las adicciones de la zona central de México.

Resultados

Las tres versiones mostraron una consistencia interna >.90, adecuada bondad de ajuste y correlaciones significativas con la Mini Entrevista Neuropsiquiátrica Internacional, v. 5.0, de criterio diagnóstico de Dependencia al Alcohol (DA). El análisis de las características operativas reveló que cada versión representó 84-85% del área bajo la curva.

Discusión y conclusión

Las tres versiones de la BEDA poseen propiedades de confiabilidad y validez para la evaluación del síndrome de dependencia al alcohol (SDA) y adecuada eficiencia diagnóstica para el mejoramiento de los pacientes con dependencia al alcohol en escenarios residenciales.

Palabras clave: Dependencia al alcohol, BEDA, propiedades psicométricas, validez de constructo, eficiencia diagnóstica.

Instituto Nacional de Psiquiatría Ramón de la Fuente Muñiz. Clinical Trials Unit on Addiction and Mental Health.

Consejo Estatal Contra las Adicciones del Estado de Hidalgo.

Correspondence: Rodrigo Marín-Navarrete, Instituto Nacional de Psiquiatría Ramón de la Fuente Muñiz, Clinical Trials Unit on Addiction and Mental Health, Calz. México-Xochimilco 101, San Lorenzo Huipulco, Tlalpan, 14370, México, D.F., México. Phone: (+52) (55) 4160 – 5481. E-mail address: rmarin@nprf.gob.mx Received first version: October 14, 2014. Accepted: June 25, 2015.

BACKGROUND

According to the literature, alcohol use disorders (AUD) are among the most important risk factors for traumatic injuries, family problems, crime, violence, and economic loss.^{1,2} They account for 4.2% of the global disability-adjusted life years (DALY) in low- and middle-income countries, as well as 6.7% of global the DALY in high-income countries.³ In Mexico, it is estimated that alcohol use is among the four leading risk factors responsible for 64.5% of the DALY in general population, which impact overall on the health care system.⁴ Therefore, it is relevant to identify individuals with AUD in order to reduce the disease burden attributable to alcohol.

Many strategies have been employed to assess AUD, including structured clinical interviews, such as the Composite International Diagnostic Interview,⁵ and semi-structured clinical interviews, such as the Mini International Neuropsychiatric Interview.⁶ Self-report questionnaires might be also useful tools to screen patients with AUD, taking less time and not requiring special training for their administration.^{7,8}

Among these, we found the Severity of Alcohol Dependence Questionnaire (SADQ), the Alcohol Dependence Scale (ADS) and the Short Alcohol Dependence Data Questionnaire (SADD), all based on the construct of the alcohol dependence syndrome (ADS) described by Edwards and Gross.9 The SADQ is a 20-item questionnaire which emphasizes withdrawal symptoms and physical dependence.¹⁰ However, it is argued that the biological emphasis of the SADQ is less likely to have discriminative power in mild and moderate dependence.¹¹ Its counterpart, the ADS, is a 25-item questionnaire that evaluates psychophysical withdrawal symptoms and loss of behavioral control.12 However, its lack of accuracy with the Diagnostic and Statistical Manual of Mental Disorders Fourth Edition (DSM-IV) alcohol dependence (AD) was recently revealed.¹³ On the other hand, the SADD, derived from the Alcohol Dependence Data (ADD) questionnaire, exhibits high correlation (r=.92) with ADD scores.14 The SADD is particularly sensitive in evaluating mild- to moderate-range dependence, because it includes cognitive, behavioral and biological components.14 The original SADD version reported item-total correlations >.50, as well as a split-half reliability of r=.87. Davidson and Raistrick¹¹ reported several studies that measured the concurrent validity of the SADD with associations from moderate to high (.49-.83). Davidson et al.¹⁵ suggested using the SADD as a 14-item questionnaire as item 9 showed low correlation with the total score.

Among the studies that evaluated the SADD in countries outside the U.K., we found studies in Brazilian and Indian populations. The first good test-retest reliability was found in a self-report of .90 and a interviewer form of .81.¹⁶ In the latter, the SADD was modified and validated with a 12-item version. Exploratory Factorial Analysis results in a unifactorial model that explained 59.3% of the variance. The authors referred that all the items demonstrated a factor loading of >.5, which is evidence of internal validity, as well as significant association among SADD scores with the amount of alcohol consumed (r=.62) and social dysfunction (r=.56) as external validity.¹⁷

In Mexico, the SADD was translated by Echeverría, Oviedo and Ayala,¹⁸ but its psychometric properties were later reported by Sánchez19 in an unpublished work using an outpatient sample. This obtained a good internal consistency (α =.84) and, in an EFA, a factorial model that explains 56.41% of the variance. The SADD is widely used as a screening instrument in a brief intervention for at-risk drinking at outpatient addiction treatment centers.²⁰ However, at residential centers, where 27.2% of the alcohol users in Mexico receive care, the psychometric properties of the SADD remain unknown.²¹ Furthermore, there is little evidence of the SADD diagnostic accuracy. Thus, the objective of the present study was to evaluate the psychometric properties and diagnostic accuracy of three SADD versions (15, 14, and 12 items) in a population receiving care at 30 addiction residential centers.

METHOD

The current study is part of a research protocol aimed at determining the psychometric properties of a set of 11 self-report clinical scales in a Mexican sample of patients with substance use disorder (SUD) at 30 addiction residential centers localized in Midwestern Mexico (in the states of México, Puebla, Querétaro and Hidalgo, and Mexico City).

Participants

For this paper, the data of the 570 participants were analyzed (figure 1). The sample comprised 502 (88.1%) men and 68 (11.9%) women. Participants were aged between 18 and 60 years, were literate and were attending a center for substance use problems, with at least one week of abstinence. Exclusion criteria included: symptoms of psychosis, mania, hypomania or cognitive impairment at the screening.

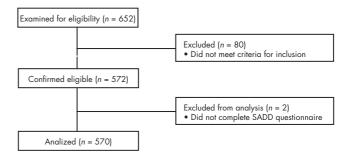


Figure 1. Flow diagram of participants.

Measures

Participants completed the Mini International Neuropsychiatric Interview in Spanish (MINI 5.0),²² a structured diagnostic interview that explores the major psychiatric disorders in the Diagnostic and Statistical Manual of Mental Disorders (DSM) and the International Classification of Diseases (ICD).²³ The MINI has shown good validity and reliability, can be administered in a relatively short time period (average 18.7 ± 11.6 min) and there is a Spanish translation of the SADD 15-item version,¹⁴ which uses a four-point, anchorage, Likert-type scale, with a minimal score of 0 and a maximal score of 45.

Procedures

All subjects were recruited for voluntary participation at each center and were assessed for eligibility utilizing the Mini-Mental State Examination (MMSE) in Spanish for dementia²⁴ and the MINI 5.0 to assess substance abuse/dependence, as well as other psychiatric conditions. A set of 11 self-report clinical scales was administered to all eligible participants immediately after screening. All interviewers were trained and certified in all study procedures. Each study procedure was approved by the Institutional Review Board of the National Institute of Psychiatry Ramón de la Fuente Muñiz.

Data analysis

Psychometric analyses included internal consistency using Cronbach's alpha and Pearson correlations. One factor model was tested using maximum likelihood CFA for three SADD versions, considering current tendencies in studies performing CFA.²⁵ The goodness-of-fit indixes employed were the following: Chi square (χ^2); degrees of freedom (df);

Table 1. Demographic characteristics of the sample

Comparative Fit Index (CFI); Root-Mean-Square Error of Approximation (RMSEA); Tucker-Lewis Index (TLI) and Goodness-of-Fit Index (GFI). To calculate diagnostic accuracy for the three SADD versions, we obtained sensitivity, specificity, Positive predictive values (PPV), Negative predictive values (NPV), multilevel likelihood ratios and efficiency values, using diagnoses obtained in the MINI 5.0 as the gold standard. In addition, Receiver Operating Characteristic (ROC) curve analysis was performed to determine the Area under the curve (AUC). Finally, DeLong's test was carried out among SADD versions to determine whether AUC differences were statistically significant.²⁶ CFA was conducted using the IBM SPSS AMOS (Analysis of Moment Structures) ver. 22.0 computer package. The remaining analyses were performed using R ver. 3.0.3 pROC software packages27 and epiR.28

RESULTS

The educational level varied significantly between males and females. Most participants reported never having been married and living in urban population (table 1). Finally, 353 males and 47 females had AD, although no statistical difference was found between both groups ($\chi^2_{(1)}$ =.041; *p*>.05).

Reliability and Validity

The mean for the three SADD scores was 17.3 (standard deviation [SD]=10.6), 15.8 (SD=10.4) and 13.8 (SD=9) (15, 14 and 12 items, respectively) in the overall sample. All SADD versions showed good internal consistency (α =.91, α =.92, and α =.91, respectively). The 15-item questionnaire exhibited significant item-total correlations >.50, with the exception of item 9, which had a low correlation (.28). The remaining SADD versions had item-total correlations >.51.

	Male (n = 502)		Female (<i>n</i> = 68)		Total (n = 570)			
	Mean or %	SD	Mean or %	SD	Mean or %	SD	Statistical difference	
Age	30.5	11	29.4	10.2	30.4	10.9	$t_{(568)} = 0.81$	
Marital status							$t_{(568)} = 0.81$ $\chi^2_{(2)} = 4.23$	
Married/Cohabitating	31.3		19.1		29.8		(-)	
Separated/Divorced	17.3		20.6		17.7			
Never been married	51.4		60.3		52.5			
Education							$\chi^2_{(3)} = 14.35*$	
No education	2.6		0.0		2.3		(0)	
Elementary/Middle school	61.6		47.1		59.8			
High school	27.8		32.4		28.4			
College or higher education	8.0		20.6		9.5			
Population							$\chi^2_{(1)} = 1.38$	
Rural	13.0		7.8		12.3		(1)	
Urban	87.0		92.2		87.7			

The goodness-of-fit measurements obtained for the 15-item model were as follows: χ^2 =269; df=81; CFI=.95; RMSEA=.06; TLI=.94, and GFI=.93. With respect to the 14-item model, we observed the following: χ^2 =264; df=70; CFI=.95; RMSEA=.07; TLI=.93, and GFI=.93. Finally, CFA for the last model (12 observed variables) suggested a model fit similar to the previous ones (χ^2 =208; df=50; CFI=.95; RMSEA=.07, TLI=.93, and GFI=.94).

Diagnostic Accuracy

The correlation between SADD and MINI 5.0 scores was .55 (p<.01). Receiver Operating Characteristic (ROC) curve analysis revealed that the 15- and 14-item questionnaires accounted for 85% of the AUC (95% confidence interval [95% CI, .81, .89]), whereas the 12-item questionnaire accounted for 84% of the AUC (95% CI, .80, .88) (figure 2).

There were three cut-off points for the SADD: a cutoff point of 14 optimized sensitivity (.80) though specificity (.81); a cut-off point of 12 optimized sensitivity (.80) though specificity (.81) and a cut-off point of 10 sensitivity (.81) specificity (.80), respectively, in the study sample. DeLong's test indicated that there were no statistically significant differences among the AUC of the 15- and 14-item SADD versions (Z=-.0668; p=.9467), 15- and 12-item versions (Z=1.1671; p=.2432), and the 14- and 12-item versions (Z=1.4007; p=.1613) (table 2).

DISCUSSION AND CONCLUSION

This study sought to evaluate three versions of the Short Alcohol Dependence Data Questionnaire (15-, 14- and 12-item models) through Confirmatory Factorial Analysis (CFA) and Receiver Operating Characteristic (ROC) curves. The three SADD versions showed good internal consistency (between .91 and .92), evidence for construct validity and a fair discriminatory diagnostic capacity. These findings prove that SADD is a single-dimension scale and supports the identification of patients with AD.

According to previous studies, item 9 (Do you try to control your drinking by giving it up completely for days or weeks at a time?) showed low item-score and within items correlations,15 which suggest that limit setting is not a common trait in people with AD. Likewise, items 13 (The morning after a heavy drinking session, do you go out of your way to avoid people?) and 15 (Do you go drinking and the next day find you have forgotten what happened the night before?) are directly associated with alcohol use-related problems rather than the ADS.^{17,29,30} However, there is little evidence on these measurement problems in the literature. Through the analysis of three SADD versions, we observed that the properties of the 12-item version barely differ from the 15- and 14-item versions, suggesting that items 9, 14 and 15 may not add significant information about the behavior of patients with alcohol use.

A CFA with a SADD one-factor solution resulted in a fair model fit for the three versions. This finding supports the Davidson et al.¹⁵ hypothesis that assumes the SADD single-factor structure. In addition, it adds to evidence pointing out the homogeneity of the ADS.³¹⁻³⁴

On the other hand, given that the dissemination of screening procedures is conducted in Mexican primary care addiction centers, the testing accuracy of screening becomes a main concern. Therefore, the determination of an optimal cut-off score for a specific population improves the practices, allowing clinicians to rule in or rule out patients into treatment with accuracy.³⁵ The diagnostic properties of the cut-off scores determined by operating characteristics curves of SADD imply that, independently of the version used, questionnaire accuracy can be trusted. However, the selection is recommended of a specific cut-off score according to the objective of the screening (whether or not higher specificity or higher sensitivity is required).

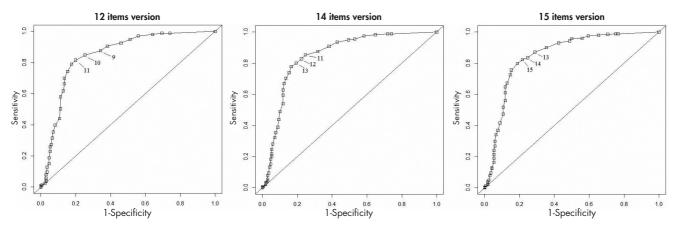


Figure 2. Receiver Operating Characteristic (ROC) curves from Short Alcohol Dependence Data Questionnaire (SADD) versions for Alcohol dependence (AD) screening.

Cut-off score (Endorsed Alcohol Dependence)	Sensitivity (95% Cl)	Specificity (95% CI)	Likelihood ratio (95% CI)	Positive Predictive value (CI 95%)	Negative Predictive value (CI 95%)	Diagnostic efficiency (95% Cl)
15 items						
11 (n = 397)	.87 (.83, .90)	.71 (.64, .78)	3.02 (2.38, 3.83)	.88 (.84, .91)	.70 (.63, .77)	.82 (.78, .85)
12 (n = 376)	.83 (.79, .87)	.75 (.68, .82)	3.38 (2.59, 4.41)	.89 (.85, .92)	.66 (.59, .73)	.81 (.77, .84)
13 (n = 366)	.82 (.78, .86)	.78 (.71, .84)	3.78 (2.83, 5.04)	.90 (.86, .93)	.65 (.58, .72)	.81 (.77, .84)
14 (n = 351)	.80 (.75, .84)	.81 (.74, .87)	4.24 (3.09, 5.81)	.91 (.87, .94)	.63 (.56, .69)	.80 (.76, .83)
15 (n = 328)	.76 (.71, .80)	.85 (.78, .90)	4.94 (3.45, 7.06)	.92 (.89, .95)	.60 (.53, .66)	.78 (.74, .81)
14 items						
9 (n= 404)	.88 (.84, .91)	.68 (.61, .75)	2.75 (2.20, 3.44)	.87 (.83, .90)	.70 (.62, .77)	.81 (.78, .84)
10 (n= 383)	.85 (.81, .89)	.75 (.68, .82)	3.45 (2.65, 4.50)	.89 (.85, .92)	.68 (.61, .75)	.82 (.78, .85)
11 (n= 369)	.83 (.79, .86)	.78 (.71, .84)	3.70 (2.79, 4.92)	.90 (.86, .93)	.66 (.59, .72)	.81 (.77, .84)
12 (n= 353)	.80 (.76, .84)	.81 (.74, .86)	4.12 (3.02, 5.62)	.91 (.87, .93)	.63 (.56, .70)	.80 (.76, .83)
13 (n= 339)	.78 (.73, .82)	.84 (.77, .89)	4.72 (3.35, 6.65)	.92 (.88, .94)	.61 (.55, .68)	.79 (.75, .82)
12 items						
7 (n= 427)	.90 (.87, .93)	.62 (.54, .69)	2.37 (1.95, 2.87)	.85 (.81, .88)	.73 (.65, .80)	.81 (.78, .85)
8 (n= 408)	.88 (.84, .91)	.66 (.58, .73)	2.56 (2.07, 3.17)	.86 (.82, .89)	.69 (.61, .76)	.81 (.77, .84)
9 (n= 382)	.85 (.81, .88)	.75 (.67, .81)	3.35 (2.58, 4.35)	.89 (.85, .92)	.68 (.60, .74)	.81 (.78, .84)
10 (n= 360)	.81 (.77, .85)	.80 (.73, .86)	4.08 (3.01, 5.52)	.91 (.87, .93)	.65 (.58, .71)	.81 (.77, .84)
11 (n= 346)	.79 (.75, .83)	.82 (.76, .88)	4.48 (3.22, 6.22)	.91 (.88, .94)	.62 (.56, .69)	.80 (.76, .83)

 Table 2. Measures of diagnostic accuracy for the three Short Alcohol Dependence Data Questionnaire (SADD) models at several of cut-off scores

In Mexico, another measure used to evaluate alcohol use is the Alcohol Use Disorders Identification Test (AUDIT). However, since this is a test designed specifically to increase the detection of a spectrum of alcohol problems, it exhibits a greater capacity to screen less severe alcohol problems, such as at-risk, harmful and hazardous drinking, and low sensitivity to dependence.³⁶⁻³⁸ Thus, SADD raises the alternative of identifying persons with AD in primary-care settings.

A limitation of this study was the small size of women sample for conducting subgroup analyses and for identifying gender-specific, optimal cut-off scores. A second limitation was the amount of self-report scales administered, which could result in biased answers as a result of respondent fatigue.

In conclusion, the three SADD versions are reliable and valid measures to assess AD. Moreover, they show optimal cut-off scores for the improvement of patient identification treatment centers and, consequently, referral to specialized care. Future studies should examine the operating characteristics of the SADD in different clinical samples (e.g., outpatients, women, psychiatric samples) to determine specific optimal cut-off points and to test the generalizability of these findings.

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Conflict of interest

No author of this paper has a conflict to interest, including specific financial interest, relationships, and/or affiliations revelant to the subject matter included in this manuscript.

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